

# SPDEs in finance and their statistical inference

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A **stochastic** partial differential equation (SPDE)

=

PDE + **random noise**

$$du(t, x) - \theta u_{xx}(t, x) dt = \sigma dW(t, x), x \in [0, 1], t \geq 0.$$

Adding noise may come from:

- ◇ passing to the limit from microscopic to mesoscopic/macroscopic level;
- ◇ model misspecification, i.e. a PDE is an approximated/ideal model;
- ◇ byproduct of other problems; e.g. filtering;
- ◇ reasonable modeling tool; e.g. finance, economics;
- ◇ other.

Note: In our terms, SPDEs  $\neq$  'random PDEs', where usually the focus is on random coefficients, random initial or/and boundary data.

## SPDEs in Finance: Interest rates modeling

Heath, Jarrow, Morton (HJM) model [HJM92] model for the instantaneous forward rate  $f(t, T) = -\partial \log P(t, T) / \partial T$ . Using Musiela parametrization  $f(t, x) := f(t, t + x)$ , it reads

$$df(t, x) = \left( \frac{\partial}{\partial x} f(t, x) + \alpha(t, x) \right) dt + \sum_{j=1}^{\infty} \sigma^j(t, x) dw^j(t),$$

where  $w^i$  are independent 1D Brownian motions, and  $\alpha$  given by no-arbitrage condition

$$\alpha(t, x) = \sum_{i=1}^{\infty} \sigma^i(t, x) \left( \int_0^{\infty} \sigma^i(t, x) du + \lambda^i(t) \right).$$

Later Brace, Gatarek, Musiela (BGM) [BGM97, Bra08]; see also [Kus00, BSC<sup>+</sup>99, SCS01], and Filipovic [Fil01].

## SPDEs in Finance: Interest rates modeling

Using PCA approach Cont [Con05], proposes to decompose

$$f(t, x) = r(t) + s(t)[y(x) + u(t, x)],$$

where  $r$  is the short rate,  $s$  the spread,  $y$  a deterministic function, and  $u$  is the deformation process, and postulated to follow the dynamics

$$du(t, x) = (\kappa u_{xx}(t, x) + u_x(t, x)) dt + \sigma W(t, x), \quad x \in [0, 1], t \geq 0,$$

where  $\kappa, \sigma > 0$  model parameters. A parabolic SPDE.

See also [BK01] for SPDEs in pricing interest rate derivatives.

More details on SPDEs in interest rates models in [CT06].

## SPDEs in Finance: forward utility measures

Musiela and Zariphopoulou [MZ10a] introduce the notion of forward utilities, in the context of optimal investment.

Assume that  $X_t^\pi$  is the wealth process corresponding to a self-financing admissible trading strategy  $\pi$ , in a standard Ito diffusion type market model. A forward utility measures  $U(t, x)$  is a function such that the optimal investment is consistent across all times,  $U(t, X_t^\pi)$  is super-martingale for any admissible  $\pi \in \mathcal{A}$ , and martingale for the optimal strategy  $\pi^*$ .

Then,  $U$  solves the following (nonlinear) SPDE

$$dU(t, x) = \frac{1}{2} \frac{|U_x(t, x)\lambda_t + \sigma_t \sigma_t^+ a(t, x)|^2}{U_{xx}(t, x)} + a(t, x) dW_t$$

See also [MZ10b, KOZ18, NZ19], and many more.

Cont and Muller [CM21]. Let  $V(t, p)$  denote the volume in LOB at time  $t$  and price  $p$ . For a small bid-ask spread and small tick size  $\delta$ , LOB can be described by a continuum approximation through its density  $v(t; p)$  representing the volume of orders per unit price:  $V(t; p) \simeq v(t; p)\delta$ . With  $S_t$  being the mid-price, define the centered order book density

$$u(t, x) = v(t, S_t + x).$$

Following some heuristics derivations, the SPDE for  $u(t, x)$  is derived

$$\begin{aligned} du &= [\eta_a \Delta u + \beta_a \nabla u + \alpha_a u + f_a] dt + \sigma_a u dW^a(t), & x \in (0, L) \\ du &= [\eta_b \Delta u + \beta_b \nabla u + \alpha_b u + f_b] dt + \sigma_b u dW^b(t), & x \in (-L, 0) \\ u(t, x) &\leq 0, \quad x < 0, & u(t, x) \geq 0, \quad x > 0 \\ u(t, 0+) &= u(t, 0-) = 0, & u(t, -L) = u(t, L) = 0. \end{aligned}$$

See also Lototsky et al. [LSZ21], Hambly et al. [HKN20].

- ◇ Implied volatility surface: Brace et al. [BFG07]
- ◇ Pricing of mortgage backed securities (MBS): Ahmad, Hambly, Ledger [AHL18]
- ◇ Large portfolio, propagation of chaos: Hambly and Kolliopoulos [HK17], and Kolli and Shkolnikov [KS18]
- ◇ Systemic risk: Hambly and Soejmark [HS19]
- ◇ DNN passing to the limit
- ◇ ...

# Statistical inference of SPDEs

Consider the stochastic heat equation

$$\begin{cases} du(t, x) - \theta u_{xx}(t, x) dt = \sigma dW(t, x), \\ u(0, x) = u_0(x), \end{cases}$$

for  $x \in G \subseteq \mathbb{R}^d$ ,  $t \geq 0$ , and where  $\sigma$  and  $\theta$  are positive parameters of interest, and  $W$  is a cylindrical Brownian motion.

## Observations:

Assume that we observe one path of the solution (or a projection of the solution) discretely or continuously in time and/or space.

## Goal:

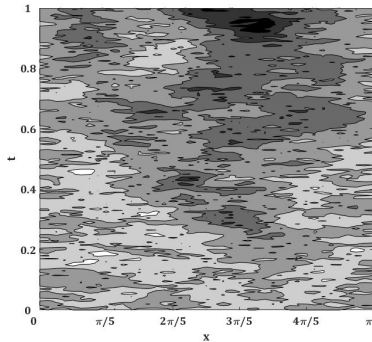
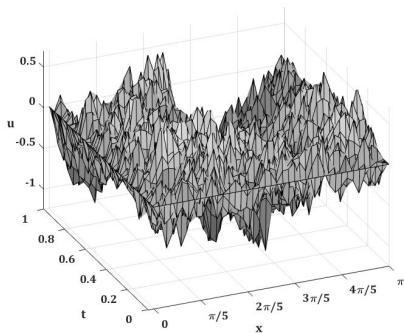
Find consistent and asymptotically normal estimators for  $\theta$  and/or  $\sigma$ .

## Why:

model estimation; model validation; ...



$$du(t, x) - \theta u_{xx}(t, x) dt = \sigma dW(t, x), \quad x \in [0, \pi]$$



**Figure:** *The observable: a sample path of the solution*

**Goal:** Find or estimate  $\theta$  and/or  $\sigma$

## Introduction: estimating drift and vol in SDEs

$dX(t) = \theta X(t)dt + \sigma X(t)dW(t)$ ,  $t \geq 0$  under probability  $\mathbb{P}$

Assume we observe one path of the solution, continuously in time.

**Girsanov Theorem (change of drift/ find likelihood ratio):**

Under some “technical conditions”,  $\exists$  a probability measure  $\mathbb{P}_0 \sim \mathbb{P}$ , such that  $dX_t = \sigma X_t dB_t$ , under  $\mathbb{P}_0$ , and

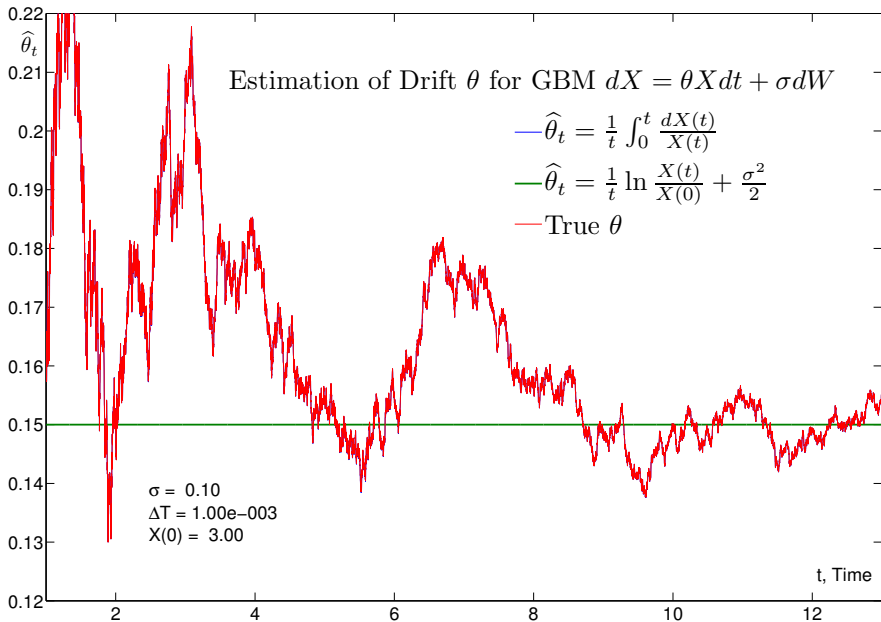
$$\frac{d\mathbb{P}}{d\mathbb{P}_0} = \exp \left( \frac{\theta}{\sigma^2} \int_0^t \frac{dX_s}{X_s} - \frac{\theta^2 t}{2\sigma^2} \right).$$

$d\mathbb{P}/d\mathbb{P}_0$  - the Likelihood Ratio (Radon-Nikodym derivative).

Maximize Likelihood Ratio  $\Rightarrow$  MLE

$$\hat{\theta}_t = \frac{1}{t} \int_0^t \frac{dX(s)}{X(s)} = \frac{1}{t} \log \frac{X(t)}{X(0)} - \frac{\sigma^2}{2},$$

$$\hat{\theta}_t \rightarrow \theta, \quad t \rightarrow \infty$$



## Intro: estimating volatility $\sigma$ in ODEs

$$dX(t) = \theta X(t)dt + \sigma X(t)dW(t), \quad t \geq 0.$$

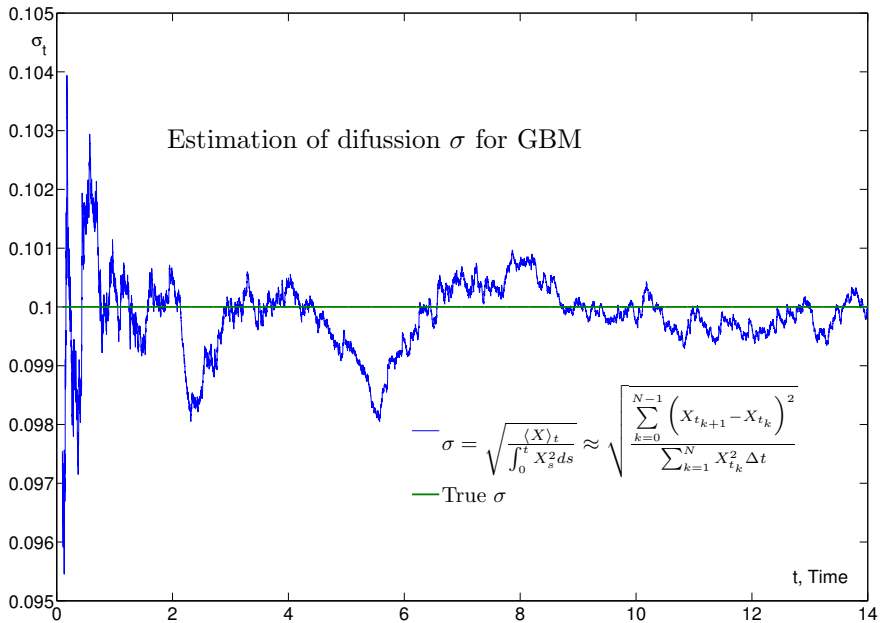
Using Quadratic Variation argument:

$$\langle X \rangle_T = \sigma^2 \int_0^T X_s^2 ds.$$

Hence,

$$\sigma = \sqrt{\frac{\langle X \rangle_T}{\int_0^T X_s^2 ds}} \approx \sqrt{\frac{\sum_{k=0}^{N-1} (X_{t_{k+1}} - X_{t_k})^2}{\sum_{k=1}^N X_{t_k}^2 \Delta t}} \rightarrow \sigma, \quad \Delta t \rightarrow 0,$$

with  $0 = t_0 < t_1 < \dots < t_N = T$ , for some fixed  $T$ .



## Stochastic ODE: conclusion

For observations  $X(t)$ ,  $t \in [0, T]$ ,  
**the drift  $\theta$  - approximated, the volatility  $\sigma$  - exactly**

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WHY?

◇ **Regular model**

1)  $\frac{d\mathbb{P}_\theta}{d\mathbb{P}_0}$  exists; 2) has a special form (LAN)

**Same statistical estimation procedure for all models**

Find MLE by maximizing likelihood ratio

◇ **Singular model** otherwise

**Individual approach**

In particular, if  $\mathbb{P}_{\sigma_1} \perp \mathbb{P}_{\sigma_2}$  for  $\sigma_1 \neq \sigma_2$ , then one may find  $\sigma$  exactly

What do we have for SPDEs?



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What do we have for SPDEs?

**Mostly singular**, hence individual approaches are needed.

- ◇ **Spectral approach:**  
sampling in the Fourier space in a continuous time setup.
- ◇ **Sampling in physical domain:**  
discrete sampling in time and/or space.
- ◇ **Local measurements:**  
sample the solution locality in space and continuously in time.

For up to date bibliography, **a dedicated web-site:**

<https://sites.google.com/view/stats4spdes/>

# Part II:

## Spectral Approach

# Spectral Approach

introduced by M. Huebner, R. Khasminskii, B. Rozovskii [HKR93, HR95]

Consider, the stochastic heat eq. with additive noise,

$$du(t, x) = \theta u_{xx}(t, x) dt + \sigma dW(t, x),$$

zero initial data, Dirichlet b.c.,  $x \in [0, \pi]$ ,  $\sigma > 0$  known,  $\theta$  unknown.

## The noise:

- ◇ the (negative) Laplace operator  $-\Delta u = -u_{xx}$  has (only) a discrete spectrum  $\nu_k = k^2, k \in \mathbb{N}$ .
- ◇ its eigenfunctions  $h_k(x) = \sqrt{2/\pi} \sin(kx)$  form a CONS in  $L^2$ .
- ◇ The noise term can be written as

$$dW(t, x) = \sum_{k=1}^{\infty} h_k(x) dw_k(t),$$

where  $w_k(t)$  are independent standard 1D Brownian motions.

# Spectral Approach

The solution can be written as

$$u(t, x) = \sum_{k=1}^{\infty} h_k(x) u_k(t),$$

where  $u_k(t) = (u, h_k)_{L^2}$  are the Fourier coefficients/modes. Clearly

$$du_k(t) + \theta k^2 u_k dt = \sigma dw_k(t), \quad k \geq 1.$$

Let  $H^N = \text{Span}\{h_1, \dots, h_N\}$ , and  $P^N$  the projection of  $H = L^2$  on  $H^N$ . Respectively, we put

$$u^N = P^N u \simeq (u_1, \dots, u_N).$$

Note that  $u^N$  follows the dynamics of a finite dimensional system of decoupled SODEs.

**The observations:** Assume that we observe one path of the first  $N$  Fourier modes continuously over a finite interval of time  $[0, T]$ , i.e. we observe/measure

$$u^N(t) = (u_1(t), \dots, u_N(t)), \quad t \in [0, T]$$

for one  $\omega \in \Omega$ .

Possible asymptotic regimes

- ◇ Large times  $T \rightarrow \infty$
- ✓ Large number of Fourier modes (fine space)  $N \rightarrow \infty$
- ◇ Small noise  $\sigma \rightarrow 0$
- ◇ Combinations of the above

Denote by  $\mathbb{P}_\theta(A) = \mathbb{P}(u^N \in A)$ ,  $A \in \mathcal{B}(C[0, T])$ , the measure generated by the solution  $u^N$ .

These are 'diagonalizable models', in the Fourier space.

Usually,  $\mathbb{P}_{\theta_1}(u^N) \sim \mathbb{P}_{\theta_2}(u^N)$ . It is a finite dimensional system of SODEs.

The likelihood ratio (the Radon–Nikodym derivative),  $\frac{d\mathbb{P}_{\theta_1}(u^N)}{d\mathbb{P}_{\theta_0}(u^N)}$  is computed by Girsanov theorem. There exists Maximum Likelihood Estimator (MLE)

$$\begin{aligned}\hat{\theta}_N &= \arg \max_{\theta_1} \log \frac{d\mathbb{P}_{\theta_1}(u^N)}{d\mathbb{P}_{\theta_0}} \\ &= - \frac{\sum_{n=1}^N \int_0^T k^2 u_k(t) du_k(t)}{\sum_{k=1}^N \int_0^T k^4 u_k^2 dt}.\end{aligned}$$

consistency & asymptotic normality

$$\hat{\theta}_N \xrightarrow[N \rightarrow \infty]{\text{a.s.}} \theta, \quad N^{\frac{3}{2}}(\hat{\theta}_N - \theta) \xrightarrow[N \rightarrow \infty]{\mathcal{D}} \mathcal{N}(0, c\sigma/T).$$

## ... since than ... see survey [Cia18]

- ◇ Bayesian: Bishwal ('02), Cheng, IgC, Gong ('18)
- ◇ Several parameters: Huebner ('97)
- ◇  $\theta(t)$ -random: Lototsky ('04)
- ◇ Small noise: Huebner ('97), Ibragimov-Khasminskii ('98,'99)
- ◇ 'almost' diagonalizable: Rozovskii-Lototsky ('97, '01)
- ◇ Additive fractional noise: IgC, Lototsky, Pospisil ('09)
- ◇ Multiplicative noise: IgC and Lototsky ('08), IgC ('10)
- ◇ Hypothesis testing: IgC and Xu ('14, '15)
- ◇ Trajectory fitting estimators: IgC, Gong, Huang ('16)
- ◇ **Nonlinear SPDE**: IgC and Glatt-Holtz ('11) 2D Navier-Stocks, semilinear SPDEs [PS20], IgC, Ruimeng Hu, Quyuan Lin [CHL23] Stochastic Primitive Equations.



## Spectral approach for nonlinear SPDEs; main ideas

On a given stochastic basis  $(\Omega, \mathcal{F}, \mathbb{F}, \mathbb{P})$ , consider the evolution equation

$$dU(t) = \theta AU(t)dt + F(U)dt + \sigma dW(t), \quad U(0) = U_0.$$

- ◇ assume that  $U(\omega, t)$  belongs to some “suitable” Hilbert space  $\mathcal{H}$ ; in particular  $U = U(\omega, t, x)$
- ◇  $(-A)$  a linear, selfadjoint, positive-defined (think  $(-\text{Laplace})^\beta$ ) in  $\mathcal{H}$  with eigenfunctions  $\{\Phi_k\}_{k \geq 1}$  CONS in  $\mathcal{H}$
- ◇  $\sigma dW(t) = \sum_{k \geq 1} \sigma_k \Phi_k dW_k(t)$ ,  $W_k, k \in \mathbb{N}$  ind. Brownian Motions
- ◇  $F$  maybe nonlinear, ‘subordinated’ to  $A$ ;  $\sigma$  known
- ◇  $U$  observed for all  $t \in [0, T]$  - **continuous observations**

### Goal:

Find estimators  $\hat{\theta}(\omega)$ ,  $\omega \in \Omega$ , for parameters  $\theta$  by **observing a single outcome**  $u = u(\omega, t) \in \mathcal{H}$  over a finite time horizon  $t \in [0, T]$ .

## Formal Procedure to Derive an Estimator

- ◇ Project the full system down to  $N$  dimensions  $P_N(\mathcal{H}) = \mathcal{H}_N \simeq \mathbb{R}^N$

$$dU^N = (\theta AU^N + \Psi_N)dt + P_N\sigma dW, \quad U(0) = U_0$$

- ◇ Let  $\mathbb{P}_\theta^{N,T}(\cdot) = \mathbb{P}(U^N \in \cdot)$  be the measure on  $C([0, T]; \mathbb{R}^N)$  generated by  $U^N$ ;

$\mathbb{P}_\theta^T$  be the measure generated by  $U$  on  $C([0, T]; \mathcal{H})$ .

- ◇ Usually (at least in linear case), we can prove that  $\mathbb{P}_{\theta_1}^{N,T} \sim \mathbb{P}_{\theta_2}^{N,T}$

Hence, get MLE type estimators  $\hat{\theta}_{N,T}$ .

- ◇ Usually (at least in linear case)  $\mathbb{P}_{\theta_1}^T \perp \mathbb{P}_{\theta_2}^T$ ;

An indication that the true parameter  $\theta$  can be found exactly.

## Formal Procedure to Derive an Estimator in Nonlinear Case

- ◇ Formally treat  $\Psi_N = P_N F(U)$  as an external and known quantity (independent of  $\theta$ )
- ◇ Assume that  $P_N \sigma$  is invertible on  $H_N$
- ◇ Take  $G := P_N \sigma(U)(P_N \sigma(U))^*$  and assume it commutes with  $A$
- ◇ For a reference values  $\theta_0$ , apply (formally) Girsanov Theorem and get the 'Likelihood Ratio' (Radon-Nikodym derivative)  $d\mathbb{P}_\theta^{N,T} / d\mathbb{P}_{\theta_0}^{N,T}$
- ◇ Maximize the Log-Likelihood Ratio  

$$\tilde{\theta}_{N,T}(\omega) := \arg \max_{\theta} d\mathbb{P}_\theta^{N,T} / d\mathbb{P}_{\theta_0}^{N,T}(\omega)$$

$$\frac{d\mathbb{P}_\theta^{N,T}}{d\mathbb{P}_{\theta_0}^{N,T}} = \exp \left[ - \int_0^T (\theta - \theta_0) \langle AU^N, GdU^N(t) \rangle \right. \\ \left. - \frac{1}{2} \int_0^T (\theta^2 - \theta_0^2) \langle AU^N, GAU^N dt \rangle \right. \\ \left. - \int_0^T (\theta - \theta_0) \langle AU^N, G\psi^N dt \rangle \right],$$

$$\hat{\theta}_{1,N} = - \frac{\int_0^T AU_N G_N dU_N + \int_0^T AU_N G_N P_N F(U) dt}{\int_0^T AU_N G_N AU_N dt},$$

$$\hat{\theta}_{2,N} = - \frac{\int_0^T AU_N G_N dU_N + \int_0^T AU_N G_N P_N F(U_N) dt}{\int_0^T AU_N G_N AU_N dt},$$

$$\hat{\theta}_{3,N} = - \frac{\int_0^T AU_N G_N dU_N}{\int_0^T AU_N G_N AU_N dt}.$$

## Idea of the proof

Easy to represent:

$$\begin{aligned}\hat{\theta}_{2,N} &= \theta + \frac{\int_0^T \langle AU^N, G \sum_{j=1}^N \sigma_j(U) \Phi_j dW_j(t) \rangle}{\int_0^T AU_N G_N AU_N dt} \\ &\quad + \frac{\int_0^T \langle AU^N, G(F^N(U) - F^N(U^N)) \rangle dt}{\int_0^T AU_N G_N AU_N dt} \\ \hat{\theta}_{3,N} &= \theta + \frac{\int_0^T \langle AU^N, G \sum_{j=1}^N \sigma_j(U) \Phi_j dW_j(t) \rangle}{\int_0^T AU_N G_N AU_N dt} \\ &\quad + \frac{\int_0^T \langle AU^N, GF^N(U^N) \rangle dt}{\int_0^T AU_N G_N AU_N dt}\end{aligned}$$

- ◇ Show that each of 'the excess term converges to zero'
- ◇ Use LLG and CLT for martingales
- ◇ sharp control of the moments of relevant parts
- ◇ 'splitting argument' to deal with the nonlinear part

## Splitting argument

Decompose  $U = \bar{U} + R = \text{linear} + \text{nonlinear}$

$$\begin{aligned}d\bar{U} &= \theta A U dt + \sigma dW, & \bar{U}(0) &= \bar{U}_0 \\dR &= \theta A R dt + F(U) dt, & R(0) &= R_0.\end{aligned}$$

- ◇ Find explicit and exact rates for the moments of the linear part
- ◇ Show that  $R$  is slightly 'more regular' than  $\bar{U}$ , and use this to show that the terms involving  $R$  vanish.

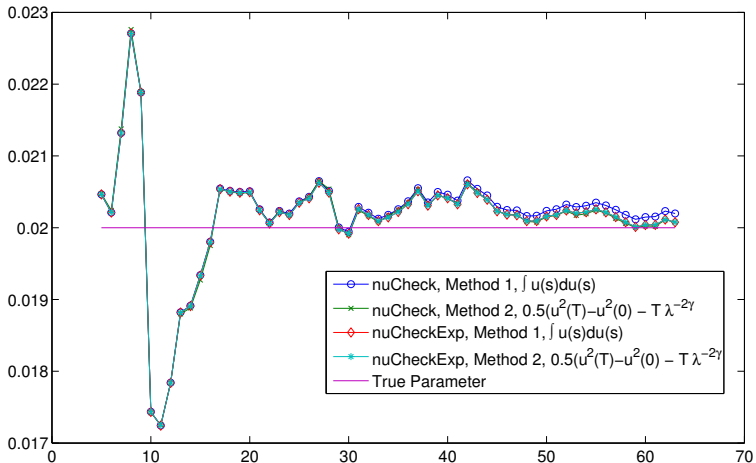
### Theorem

*All three estimators  $\hat{\theta}_{1,N}, \hat{\theta}_{2,N}, \hat{\theta}_{3,N}$ , are consistent and asymptotically normal,*

$$\hat{\theta}_{k,N} \xrightarrow[N \rightarrow \infty]{a.s.} \theta, \quad N^J (\hat{\theta}_{k,N} - \theta) \xrightarrow[N \rightarrow \infty]{\mathcal{D}} \mathcal{N}(0, \sigma^*),$$

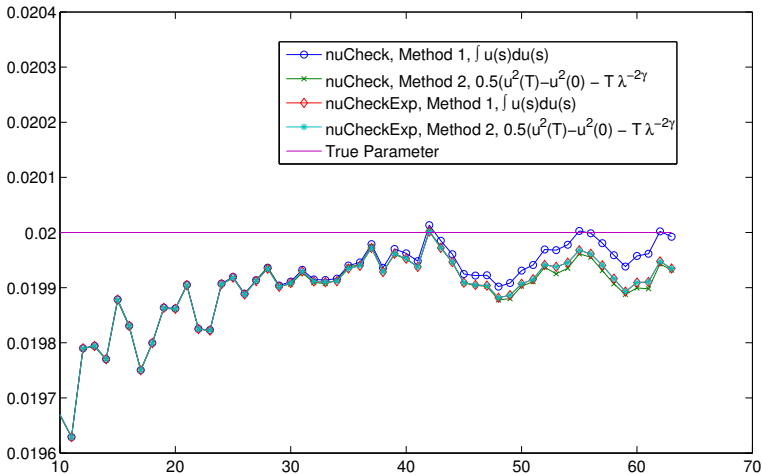
*for some known rate of convergence  $J$ , depending on the dimension of the space and order of  $A$ , and with a known asymptotic variance  $\sigma^*$ .*

Heat Equation  $du = \nu u_{xx} dt + \sigma^{-\gamma} dW$ ,  $\nu=0.02$ ,  $T=1$ ,  $\alpha=0.1$ ,  $\gamma= 0.1$



$$du = \nu u_{xx} dt + dW(t, x), \quad u(0, x) = u_0, \quad u(t, 0) = u(t, 1) = 0$$

Burgers Equation  $du = \nu u_{xx} dt - \beta u \cdot u_x dt + \sigma^{-\gamma} dW$ ,  $\nu=0.02$ ,  $T=1$ ,  $\alpha=0.1$ ,  $\gamma=0.1$



$$du = \nu u_{xx} dt - uu_x dt + dW(t, x), \quad u(0, x) = u_0, \quad u(t, 0) = u(t, 1) = 0$$



## Part III:

Discrete sampled SPDEs in physical domain

## Stats for SPDEs: discrete sampling in physical domain

The parameter estimation problem for (linear) SPDEs when the solution is discretely sampled in space and/or time component was addressed systematically only recently by quite different methods: [CH20, BT20] and consequently in [BT19, Cho20, CDVK20, Cho19, KU21, KT19a, KT19b, HT21, CK22, SST20, CKP23], and to [PR97, PT07] for earlier studies, as well as the recent work [HT23] on reaction-diffusion equations.

**Our approach:** use the exact order of regularity and correctly chosen power variation.

[CKP23] C., H.-J. Kim and G. Pasemann *Statistical analysis of discretely sampled semilinear SPDEs: a power variation approach*, forthcoming in *Stochastics and Partial Differential Equations: Analysis and Computations*, 2023+.

[CK22] C. and H.-J. Kim, *Parameter estimation for discretely sampled stochastic heat equation driven by space-only noise revised*, SPA, 22, pp. 1-30, 2022.

[CH20] C. and Y. Huang, *A note on parameter estimation for discretely sampled SPDEs* *Stochastics and Dynamics* 20(3), pp. 2050016, 2020 (28 pages, preprint 2017).

## The main goal

Statistical analysis of **discretely sampled** (SPDEs) of the form

$$dX_t(x) = -\theta(-\Delta)^{\alpha/2}X_t(x)dt + F(X_t(x))dt + \sigma(-\Delta)^{-\gamma}dW_t(x),$$

for  $x \in [0, 1]$ ,  $t > 0$ , Dirichlet boundary conditions and zero initial data, and where

- ◇  $\alpha > 0, \gamma \geq 0$  are given constants,
- ◇  $W$  is a cylindrical Wiener process on  $L^2([0, 1])$ ,
- ◇  $F$  is a (nonlinear) operator acting on some appropriate Hilbert space,
- ◇  $\theta, \sigma$  are the **parameters of interest** (unknown).

Simplest nontrivial example: Stochastic heat equation

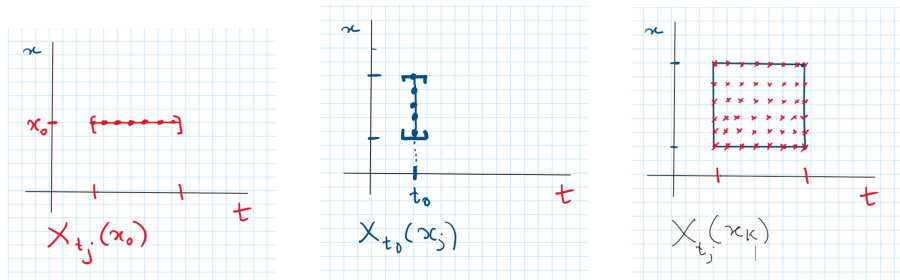
$$du(t, x) = \theta u_{xx}(t, x)dt + \sigma \sum_{k \geq 1} k^{-2\gamma} \sin(k\pi x) dw_k(t), \quad x \in [0, 1],$$

observed at some discrete points  $(t_k, x_j)$ .

## What is the problem?

$$dX_t(x) = -\theta(-\Delta)^{\alpha/2} X_t(x) dt + F(X_t(x)) dt + \sigma(-\Delta)^{-\gamma} dW_t(x)$$

- ◇ We know how to treat the case of non-smooth path, both in time and space; for  $\gamma = 0$ ,  $X \in C_{t,x}^{1/4-, 1/2-}$  [CH20].
- ◇ It is enough to sample discretely the solution  $X(t, x)$  in space and/or time at one time point, or one space point, or on a space-time mesh.



**Figure:** *Sampling schemes*

$$dX_t(x) = -\theta(-\Delta)^{\alpha/2} X_t(x) dt + F(X_t(x)) dt + \sigma(-\Delta)^{-\gamma} dW_t(x)$$

- ◇ Larger  $\gamma$  gives smoother solutions.
- ◇ Regularity in  $t$  can't get better than Hölder  $1/2$ -, and the existing power variation methods apply.
- ◇  $X(t, \cdot)$  can reach any order of smoothness, when  $\gamma \nearrow \infty$ .  
For example, when  $F = 0$ ,  $X(t, \cdot)$  is Hölder continuous of order  $2\gamma + \alpha/2 - 1/2$ .
- ◇ We focus on sampling the solution  $X$  discretely in spatial variable  $x$  and for a fixed  $t > 0$ .

## Main line of reasoning

- ◇ Take the maximal number of (classical) derivatives in  $x$ , say

$$m := \lfloor 2\gamma + \alpha/2 - 1/2 \rfloor$$

- ◇ Expect that

$$\partial_x^m X_t(x) \sim \text{fBM}^H + \text{'smooth process'},$$

with  $H = 2\gamma + \alpha/2 - 1/2 - m$ .

- ◇ Adapt the existing results on power variations from [CH20, KT19a, KT19b].
- ◇ **It works!**

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- ◇ Adapt the existing results on power variations from [CH20, KT19a, KT19b].
- ◇ **It works!**

**However**, assuming that the process  $\partial_x^m X_t(x)$ ,  $x \in (0, 1)$  is observed, practically speaking is unrealistic.

## Main line of reasoning

- ◇ **Natural idea:** approximate the derivatives  $\partial_x^m X_t(x)$  by using the discrete measurements of the solution itself, for example by finite differences.
- ◇ **It does not work!**

Such approximations typically yield a nontrivial and non-vanishing bias in the estimators; see also [CKL20], [CK22].

## Main results

Find the needed adjustments (biases) for the naively approximated estimators and prove consistency and asymptotic normality.



## Notations

For real valued measurable function  $X_t, t \in \mathbb{R}$ , we put

$$JX_t := \int_0^t X_r \, dr, \quad t \in \mathbb{R},$$

$$\Delta_h X_t := X_{t+h} - X_t, \quad t \in \mathbb{R}, \quad h > 0.$$

For  $M, m \in \mathbb{N}_0 := \mathbb{N} \cup \{0\}$ , we have

$$\Delta_h^M (J^m X_t) = \sum_{k=0}^M (-1)^{M-k} \binom{M}{k} J^m X_{t+kh}, \quad t \in \mathbb{R}, \quad h > 0.$$

$C(\mathbb{R})$  denotes the space of continuous and bounded functions on  $\mathbb{R}$ , endowed with  $\|f\|_\infty := \sup |f|$ .

$C^k(\mathbb{R}) := \{f \in C(\mathbb{R}) : \|f\|_{C^k(\mathbb{R})} := \sum_{j \leq k} \|D^j f\|_\infty < \infty\}$ ,  
for  $k \in \mathbb{N}$ , and with  $D$  being the differential operator.

- ◇ Let  $\pi = \{t_0, \dots, t_N\}$  be the uniform partition of size  $N$  of the interval  $[a, b] \subset [0, T]$
- ◇ Put  $h := h_N := (b - a)/N = t_{k+1} - t_k, k = 0, \dots, N$ .
- ◇ For fixed  $s > 0, q, M, N \in \mathbb{N}$ , such that  $N > M$ , we define

$$V_{q,M,s,N}(X) := \frac{1}{b-a} \sum_{k=0}^{N-M} h \left| \frac{\Delta_h^M X_{t_k}}{h^s} \right|^q.$$

- ◇ The  **$\Delta$ -power variation of order  $(q, M, s)$  of process  $X$**  is defined as

$$V_{q,M,s}(X) := \mathbb{P} - \lim_{N \rightarrow \infty} V_{q,M,s,N}(X),$$

provided that the limit (in probability) exists.

Note that  $V_{p,1,1}$  corresponds to the (normalized) power variation of order  $p$ .

## Theorem

Let  $q \geq 1$ ,  $s > 0$ ,  $M \in \mathbb{N}$  with  $M > s$ . Assume that  $X \in C^s([a, b])$  and for some  $\alpha > 0$ ,  $\Sigma \geq 0$ , the following limit exists

$$h_N^{-\alpha} (V_{q,M,s,N}(X) - V_{q,M,s}(X)) \xrightarrow{d} \mathcal{N}(0, \Sigma), \quad \text{as } N \rightarrow \infty,$$

where  $\mathcal{N}(0, \Sigma)$  is a Gaussian random variable with mean zero and variance  $\Sigma$ . Then, for any  $Y \in C^{s+\eta}([a, b])$  with  $\eta > \alpha$ , and  $M > s + \alpha$ ,

$$h_N^{-\alpha} (V_{q,M,s,N}(X + Y) - V_{q,M,s}(X)) \xrightarrow{d} \mathcal{N}(0, \Sigma), \quad \text{as } N \rightarrow \infty.$$

Proof based on equivalence of Hölder-Zygmund spaces and classical Hölder spaces.

## The case of fBM

- ◇ A fractional Brownian motion (fBm) with Hurst index  $H \in (0, 1)$  is a centered Gaussian process  $B^H = (B_t^H)_{t \in \mathbb{R}}$  such that

$$\mathbb{E}(B_t^H B_r^H) = \frac{1}{2} (|t|^{2H} + |r|^{2H} - |t - r|^{2H}), \quad t, r \in \mathbb{R}.$$

- ◇ A fBM  $B^H$  is  $H$ -self-similar process with stationary increments.
- ◇ Generally speaking, differences of integrals of fBm are not self-similar in the usual sense.
- ◇ We extend the notion of self-similarity to a parameterized family of processes, say  $X^{(h)}$ ,  $h > 0$ . We say that  $X^{(h)}$  is **parameterized  $s$ -self-similar** if the law of  $(h^{-s} X_{ht}^{(h)})_{t \in \mathbb{R}}$  is independent of  $h > 0$ .

If  $M \geq m$ , then  $\Delta_h^M J^m B^H$  is parameterized  $(m + H)$ -self-similar and has stationary increments.

# Main result for fBM

## Theorem

Let  $M > m \geq 0$  and  $q \geq 1$  be integers,  $s = H + m$ , and assume that either of the following assumptions is satisfied:

- (i)  $M = m + 1$  and  $0 < H < 3/4$ ,
- (ii)  $M \geq m + 2$  and  $0 < H < 1$ .

Then, there exists  $\sigma_{q,M,s} > 0$  such that

$$\sqrt{N} \left( V_{q,M,s,N} (J^m B^H) - \tau_q \mu_{M,s}^{q/2} \right) \xrightarrow{d} \mathcal{N} \left( 0, \sigma_{q,M,s}^2 \mu_{M,s}^q \right), \quad \text{as } N \rightarrow \infty,$$

where  $\tau_q := \mathbb{E}|Z|^q$  with  $Z \sim \mathcal{N}(0, 1)$ .

Moreover, if  $q$  is an even number, then  $\sigma_{q,M,s}^2 = \sum_{k=1}^q \binom{q}{k}^2 \tau_{q-k}^2 \rho_{k,M,s}^2$ ,

where  $\rho_{q,M,s}^2 := q! \sum_{\ell \in \mathbb{Z}} (\rho_{M,s}(\ell))^q$ .

Recall that  $V_{q,M,s}(X) := \mathbb{P} - \lim_{N \rightarrow \infty} \frac{1}{b-a} \sum_{k=0}^{N-M} h \left| \frac{\Delta_h^M X_{t_k}}{h^s} \right|^q$ .

## Remarks

- ◇ The proof is based on a version of Breuer-Major Theorem.
- ◇ The limit of  $V_{q,M,s,N}(J^m B^H)$  depends through  $\mu_{M,s}$  on the regularity  $s$  of the process as well as the number of differences  $M$ .
- ◇ Even for small  $h$  it is not possible to approximate the rescaled finite difference operator  $h^{-1}\Delta_h$  in the definition of  $V_{q,M,s,N}(J^m B^H)$  by a derivative operator without introducing a non-trivial bias, due to the change in  $M$  and  $s$ .
- ◇ The constant  $\mu_{M,s}$  can be easily computed.

If  $M = 1$ ,  $m = 0$  and  $0 < H < 3/4$ , then  $\mu_{M,s} = 1$ .

If  $M = 2$ ,  $m = 1$  and  $H = 1/4$ , then

$$\mu_{M,s} = (\sqrt{2} - 1) \frac{16}{15} \approx 0.44.$$

If  $M = 2$ ,  $m = 1$  and  $H = 1/2$ , then  $\mu_{M,s} = 2/3$ .

## The case of semilinear SPDEs

- ◇ We consider SPDEs on  $\mathcal{D} = (0, 1)$  with zero boundary conditions.
- ◇ Set  $\Phi_k(x) = \sqrt{2} \sin(k\pi x)$  and  $\lambda_k = k^2\pi^2$ ,  $k \in \mathbb{N}$ . These are the eigenfunctions and eigenvalues of the Laplacian  $\Delta = \partial_{xx}$ .
- ◇ The set  $\{\Phi_k\}_{k \in \mathbb{N}}$  forms an orthonormal basis in  $L^2(\mathcal{D})$ .
- ◇ Put  $H^s(\mathcal{D}) := \{u \in L^2 \mid \sum_{k=1}^{\infty} \lambda_k^s(u, \Phi_k)^2 < \infty\}$ , for  $s \in \mathbb{R}$ ,

We consider the following semilinear SPDE on  $L^2(\mathcal{D})$ :

$$dX_t = \left( -\theta(-\Delta)^{\alpha/2} X_t + F(X_t) \right) dt + \sigma(-\Delta)^{-\gamma} dW_t, \quad X_0 \in L^2(\mathcal{D}),$$

where  $\alpha, \theta, \sigma > 0$ ,  $W$  is a cylindrical Wiener process on  $L^2(\mathcal{D})$ ,  $\gamma > 1/4 - \alpha/4$ , and  $F$  is a nonlinear operator.

We assume that the above SPDE is well-posed in  $L^2(\mathcal{D})$  in the analytically mild and probabilistically weak sense.

## Splitting of the solution argument

We use the splitting of the solution argument (see [CGH11, PS20, ACP20]), by writing  $X = \bar{X} + \tilde{X}$ , where

$$\begin{aligned}d\bar{X}_t &= -\theta(-\Delta)^{\alpha/2}\bar{X}_t dt + \sigma B dW_t, & \bar{X}_0 &= 0, \\d\tilde{X}_t &= \left(-\theta(-\Delta)^{\alpha/2}\tilde{X}_t + F(\bar{X}_t + \tilde{X}_t)\right) dt, & \tilde{X}_0 &= X_0.\end{aligned}$$

The solution to the linear equation can be written as a Fourier series

$$\begin{aligned}\bar{X}_t &= \sigma \int_0^t e^{-\theta(t-r)(-\Delta)^{\alpha/2}} B dW_r = \sum_{k=1}^{\infty} \left( \sigma \lambda_k^{-\gamma} \int_0^t e^{-\theta(t-r)\lambda_k^{\alpha/2}} dW_r^{(k)} \right) \Phi_k \\ &=: \sum_{k=1}^{\infty} \bar{x}_k(t) \Phi_k,\end{aligned}$$



## Fine continuity properties

### Proposition

For any  $s < 2\gamma + \alpha/2 - 1/2$ , it holds that  $\bar{X} \in C(0, T; C^s(\mathcal{D}))$ .

Hence  $\bar{X}$  has up to  $\lfloor 2\gamma + \alpha/2 - 1/2 \rfloor =: m$  classical derivatives. We call  $s^* = 2\gamma + \alpha/2 - 1/2$  the **optimal regularity**. We assume that  $s^* \notin \mathbb{N}$ .

### Proposition

Assume that there exist  $\eta, \epsilon > 0$ ,  $0 \leq s_0 < s^*$ , and a continuous function  $g : [0, \infty) \rightarrow [0, \infty)$ , such that for any  $s_0 \leq s < s^*$ ,

$$\|F(u)\|_{s+\eta-\alpha+\epsilon} \leq g(\|u\|_s),$$

where, as before,  $\|\cdot\|_s$  denotes the Hölder-Zygmund norm. Let  $X \in C(0, T; C^{s_0}(\mathcal{D}))$  and  $X_0 \in C^{s^*+\eta}(\mathcal{D})$ . Then we have  $\tilde{X} \in C(0, T; C^{s+\eta}(\mathcal{D}))$ , for any  $0 \leq s < s^*$ , and  $X \in C(0, T; C^s(\mathcal{D}))$ .

## Examples

The above conditions are satisfied for large classes of SPDEs. Some important examples of the nonlinearity  $F$ :

- 1) *(fractional) Heat equation*: In the case  $F = 0$ , the SPDE becomes linear, sometimes called fractional heat equation, and the Lip condition is trivially satisfied for any  $\eta > 0$ .
- 2) *Reaction-diffusion equation*: Let  $F(u)(x) = f(u(x))$ , where  $f$  is a polynomial function or  $f \in C_b^\infty(\mathbb{R})$ . Then Lip condition is true for any  $0 < \eta < 2$ .
- 3) *Advection-diffusion equation*: Let  $F(u) = v\partial_x u$  for a given  $v \in C^\infty(\mathcal{D})$ . Then Lip condition holds with any  $0 < \eta < 1$ .
- 4) If  $F = F_1 + F_2$ , for some  $F_1, F_2$  that satisfy Lip condition with continuous functions  $g_1, g_2$ , then  $F$  satisfies Lip condition with  $g = g_1 + g_2$ .

# Parameter estimation for SPDEs

## Theorem

Let  $t > 0$ ,  $m \in \mathbb{N}_0$ ,  $0 < H < 1$  such that  $m + H = s^* = 2\gamma + \alpha/2 - 1/2$ . Let  $M, q \in \mathbb{N}$ , and assume that either  $M = m + 1$  with  $H < 1/2$  or  $M \geq m + 2$ . Suppose that Lip condition holds for some  $\eta > 1/2$ , and assume that  $\theta$  is known. Then

$$\widehat{\sigma}_N^{q,M} := \tau_q^{-1} (2\theta / (\nu_H \mu_{M,s^*}))^{q/2} V_{q,M,s^*,N}(X_t)$$

with  $\nu_H := -\frac{2}{\pi} \Gamma(-2H) \cos(\pi H)$ , is a consistent estimator for  $\sigma^q$ , and for any  $\epsilon > 0$ ,

$$\widehat{\sigma}_N^{q,M} = \sigma^q + o_{\mathbb{P}}(N^{-1/2+\epsilon}).$$

If  $s^* \in 1/2 + \mathbb{N}_0$ , then also

$$\sqrt{N} \left( \widehat{\sigma}_N^{q,M} - \sigma^q \right) \xrightarrow{d} \mathcal{N} \left( 0, \frac{\sigma^{2q}}{\tau_q^2} \sigma_{q,M,s^*}^2 \right), \quad \text{as } N \rightarrow \infty.$$

Similar statement holds true for parameter  $\theta$ .

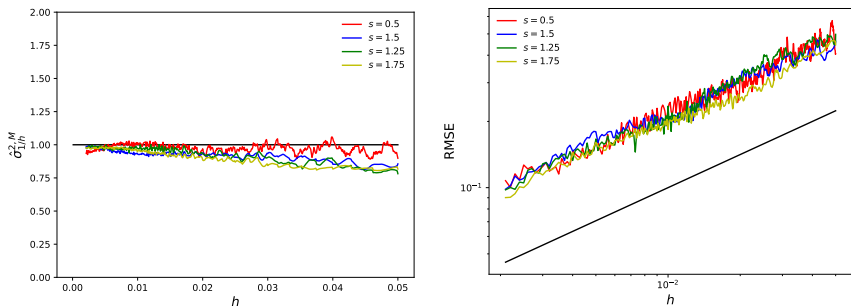
## Numerical example

Consider the stochastic heat equation

$$dX_t = \theta \Delta X_t dt + \sigma (-\Delta)^{-\gamma} dW_t,$$

with initial condition  $X_0 = 0$  on  $\mathcal{D} = [0, 1]$  with Dirichlet BC.

- ◇ Take true values of the parameters  $\theta, \sigma = 1$ .
- ◇ The smoothing parameter  $\gamma \in \{0.0, 0.375, 0.5, 0.625\}$ , which correspond to the regularity level  $s^* = 2\gamma + 1/2 \in \{0.5, 1.25, 1.5, 1.75\}$ .
- ◇ Simulate the path using the Fourier series decomposition of the solution by taking  $N_0 = 1 \times 10^4$  eigenmodes.
- ◇ Eigenmodes are simulated by the Euler implicit scheme with  $\delta t = 1 \times 10^{-8}$ .
- ◇ The solution is computed at  $N_0 + 1$  uniformly spaced spatial grid points with step size  $h = 1 \times 10^{-4}$ .
- ◇ Assume that the solution  $X$  is observed at time  $T = 1$  on spatial grid points belonging to the interval  $[a, b]$ , with  $a = 0.2, b = 0.8$ .
- ◇ Apply the main Theorem, with  $q = 2$  and  $M = \lceil s^* \rceil + 2$ .



**Figure:** Estimation of  $\sigma$ .

Left panel: the average of 100 Monte Carlo estimates as function of spatial sampling resolution  $h$ . The solid black line corresponds to the true value  $\sigma = 1.0$ .

Right panel: The RMSE (root mean square error) as function of  $h$ . The black line corresponds to the theoretical convergence rate  $h^{1/2}$ .

**Thank You !**

The end of the talk  
but not of the story.

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## Appendix

Breuer-Major Theorem [NP12, Theorem 7.2.4 ].

### Theorem

Let  $Y = \{Y_k\}_{k \in \mathbb{Z}}$  be a centered stationary Gaussian sequence with unit variance, and  $f(x) = \sum_{q=d}^{\infty} a_q H_q(x)$ ,  $a_q \in \mathbb{R}$ , where  $H_q$  is the  $q$ -th Hermite polynomial. Assume that

$$\sum_{\ell \in \mathbb{Z}} |\rho(\ell)|^d < \infty, \quad (3.1)$$

where  $\rho(\ell) = \mathbb{E}(Y_0 Y_\ell)$ ,  $\ell \in \mathbb{Z}$ . Then,

$$\mathbf{w}\text{-}\lim_{N \rightarrow \infty} \frac{1}{\sqrt{N}} \sum_{k=1}^N f(Y_k) = \mathcal{N} \left( 0, \sum_{q=d}^{\infty} q! a_q^2 \sum_{\ell \in \mathbb{Z}} \rho(\ell)^q \right).$$

## The noise $W(t, x)$

$W(t)$  is a cylindrical Brownian motion, if it is an  $\mathcal{H} = L^2(\Lambda)$ -valued continuous Gaussian process with  $W(0) = 0$ , and *covariance structure*

$$\mathbb{E}[\langle W(t), g \rangle_\Lambda \cdot \langle W(s), f \rangle_\Lambda] = \min(t, s) \cdot \langle f, g \rangle_\Lambda$$

$\dot{W}(t, x)$ ,  $t \geq 0, x \in \Lambda \subset \mathbb{R}^d$  is called *space-time white noise*:

- ◇ a zero mean Gaussian process with covariance

$$\mathbb{E}[\dot{W}(t, x)\dot{W}(s, y)] = \delta(t - s)\delta(x - y)$$

- ◇  $\dot{W}(t, x) = \sum_{k \geq 1} h_k(x)\dot{w}_k(t)$ ,  
where  $\{h_k(x)\}_{k \in \mathbb{N}}$  is an orthonormal basis in  $L^2(\Lambda)$ , and  $\{\dot{w}_k\}_{k \in \mathbb{N}}$  are independent 1d white-noises.

- ◇ a random generalized function on  $L_2((0, T) \times \Lambda)$

$$\dot{W}(t, x) = \int_0^T \int_\Lambda f(t, x) dW(t, x).$$